

Homework 8

Solutions!

1. Let $X \sim \Gamma(\alpha, \beta)$. Derive the variance of X without using the Moment Generating Function of X

Solution:

Proof.

$$\begin{aligned}V[X] &= E[X^2] - E^2[X] \\&= E[X^2] - (\alpha\beta)^2 \\E[X^2] &= \int_{-\infty}^{\infty} x^2 f_X(x) dx \\&= \int_{-\infty}^0 x^2 f_X(x) dx + \int_0^{\infty} x^2 f_X(x) dx \\&= 0 + \int_0^{\infty} x^2 \frac{1}{\Gamma(\alpha)\beta^\alpha} x^{\alpha-1} e^{-\frac{x}{\beta}} dx \\&= 0 + \int_0^{\infty} x \frac{1}{\Gamma(\alpha)\beta^\alpha} x^\alpha e^{-\frac{x}{\beta}} dx \\&= \int_0^{\infty} x \frac{1}{\Gamma(\alpha)} \left(\frac{x}{\beta}\right)^\alpha e^{-\frac{x}{\beta}} dx\end{aligned}$$

Integration by substitution

$$u = \frac{x}{\beta} \Rightarrow x = \beta u$$

$$\Rightarrow du = \frac{dx}{\beta}$$

$$\Rightarrow \beta du = dx$$

$\Rightarrow u$ Ranges from 0 to ∞

When x ranges from 0 to ∞

$$\begin{aligned}\Rightarrow \int_0^{\infty} x \frac{1}{\Gamma(\alpha)} \left(\frac{x}{\beta}\right)^\alpha e^{-\frac{x}{\beta}} dx &= \frac{1}{\Gamma(\alpha)} \int_0^{\infty} (\beta u)(u)^\alpha e^{-u} \beta du \\&= \frac{\beta^2}{\Gamma(\alpha)} \int_0^{\infty} (u)^{\alpha+1} e^{-u} du \\&= \frac{\beta^2}{\Gamma(\alpha)} \int_0^{\infty} (u)^{(\alpha+2)-1} e^{-u} du\end{aligned}$$

$$\begin{aligned}
&= \frac{\beta^2}{\Gamma(\alpha)}\Gamma(\alpha + 2) \\
&= \frac{\beta^2}{\Gamma(\alpha)}(\alpha + 1)\Gamma(\alpha + 1) \\
&= \frac{\beta^2}{\Gamma(\alpha)}(\alpha + 1)(\alpha)\Gamma(\alpha) \\
&= \alpha^2\beta^2 + \alpha\beta^2 \\
\Rightarrow V[X] &= E[X^2] - E^2[X] \\
&= \alpha^2\beta^2 + \alpha\beta^2 - (\alpha\beta)^2 \\
&= \alpha\beta^2
\end{aligned}$$

□

2. Let $X \sim \text{Beta}(\alpha, \beta)$. Derive the variance of X .

Solution:

Proof.

$$\begin{aligned}
V[X] &= E[X^2] - E^2[X] \\
&= E[X^2] - \left(\frac{\alpha}{\alpha + \beta}\right)^2 \\
E[X^2] &= \int_{-\infty}^{\infty} x^2 f_X(x) dx \\
&= \int_{-\infty}^0 x f_X(x) dx + \int_0^1 x^2 f_X(x) dx + \int_1^{\infty} x f_X(x) dx \\
&= 0 + \int_0^1 x^2 \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} x^{\alpha-1} (1-x)^{\beta-1} dx + 0 \\
&= \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} \int_0^1 x^{\alpha+1} (1-x)^{\beta-1} dx \\
&= \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} \int_0^1 x^{(\alpha+2)-1} (1-x)^{\beta-1} dx \\
&= \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} B(\alpha + 2, \beta) \leftarrow \text{Definition of Beta function} \\
&= \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} \frac{\Gamma(\alpha + 2)\Gamma(\beta)}{\Gamma(\alpha + \beta + 2)} \leftarrow \text{Beta function Theorem} \\
&= \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} \frac{(\alpha + 1)\Gamma(\alpha + 1)\Gamma(\beta)}{(\alpha + \beta + 1)\Gamma(\alpha + \beta + 1)} \\
&= \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} \frac{(\alpha + 1)\alpha\Gamma(\alpha)\Gamma(\beta)}{(\alpha + \beta + 1)(\alpha + \beta)\Gamma(\alpha + \beta)}
\end{aligned}$$

$$\begin{aligned}
&= \frac{\Gamma(\alpha + \beta) \Gamma(\alpha)\Gamma(\beta)}{\Gamma(\alpha)\Gamma(\beta) \Gamma(\alpha + \beta)} \frac{(\alpha + 1)\alpha}{(\alpha + \beta + 1)(\alpha + \beta)} \\
&= \frac{\alpha^2 + \alpha}{(\alpha + \beta)(\alpha + \beta + 1)} \\
\Rightarrow V[X] &= E[X^2] - E^2[X] \\
&= \frac{\alpha^2 + \alpha}{(\alpha + \beta)(\alpha + \beta + 1)} - \left(\frac{\alpha}{\alpha + \beta}\right)^2 \\
&= \frac{(\alpha^2 + \alpha)(\alpha + \beta)}{(\alpha + \beta)^2(\alpha + \beta + 1)} - \frac{\alpha^2(\alpha + \beta + 1)}{(\alpha + \beta)^2(\alpha + \beta + 1)} \\
&= \frac{(\alpha^2 + \alpha)(\alpha + \beta) - \alpha^2(\alpha + \beta + 1)}{(\alpha + \beta)^2(\alpha + \beta + 1)} \\
&= \frac{\alpha^3 + \alpha^2 + \alpha^2\beta + \alpha\beta - \alpha^3 - \alpha^2\beta - \alpha^2}{(\alpha + \beta)^2(\alpha + \beta + 1)} \\
&= \frac{\alpha\beta}{(\alpha + \beta)^2(\alpha + \beta + 1)}
\end{aligned}$$

□

3. Let $X \sim Exp(\delta)$. Show tha the CDF of x , $F_X(x)$ is

$$F_X(x) = \begin{cases} 1 - e^{-\frac{x}{\delta}} & 0 < x \\ 0 & \text{else} \end{cases}$$

Solution:

Proof.

$$\begin{aligned}
F_X(x) &= \int_{-\infty}^x f_X(t)dt \\
&= \begin{cases} \int_{-\infty}^x 0dt & x < 0 \\ \int_{-\infty}^0 0dt + \int_0^x \frac{1}{\delta}e^{-\frac{t}{\delta}}dt & 0 < x \end{cases} \\
&= \begin{cases} 0 & x < 0 \\ 0 + -e^{-\frac{t}{\delta}}|_0^x & 0 < x \end{cases} \\
&= \begin{cases} 0 & x < 0 \\ -e^{-\frac{x}{\delta}} - (-e^{-\frac{0}{\delta}}) & 0 < x \end{cases} \\
&= \begin{cases} 0 & x < 0 \\ 1 - e^{-\frac{x}{\delta}} & 0 < x \end{cases} \\
&= \begin{cases} 1 - e^{-\frac{x}{\delta}} & 0 < x \\ 0 & \text{else} \end{cases}
\end{aligned}$$

□

4. Let $X \sim \Gamma(\alpha, \beta)$ and let r be a real constant such that $r > -\alpha$. Prove that

$$E[X^r] = \frac{\beta^r \Gamma(\alpha + r)}{\Gamma(\alpha)}$$

Solution:

Proof.

$$\begin{aligned} E[X^r] &= \int_{-\infty}^{\infty} x^r f_X(x) dx \\ &= \int_0^{\infty} x^r \frac{1}{\Gamma(\alpha)\beta^\alpha} x^{\alpha-1} e^{-\frac{x}{\beta}} dx \\ &= \frac{1}{\Gamma(\alpha)\beta^\alpha} \int_0^{\infty} x^{\alpha+r-1} e^{-\frac{x}{\beta}} dx \\ &= \frac{1}{\beta^\alpha} \frac{\Gamma(\alpha+r)\beta^{\alpha+r}}{\Gamma(\alpha+r)\beta^{\alpha+r}} \int_0^{\infty} x^{\alpha+r-1} e^{-\frac{x}{\beta}} dx \\ &= \frac{\Gamma(\alpha+r)\beta^{\alpha+r}}{\Gamma(\alpha)\beta^\alpha} \int_0^{\infty} \frac{1}{\Gamma(\alpha+r)\beta^{\alpha+r}} x^{\alpha+r-1} e^{-\frac{x}{\beta}} dx \\ &\quad \text{This is the (valid) pdf of } X \sim \Gamma(\alpha+r, \beta) \text{ since } r > -\alpha \uparrow \\ &= \frac{\Gamma(\alpha+r)\beta^{\alpha+r}}{\Gamma(\alpha)\beta^\alpha} \\ &= \frac{\Gamma(\alpha+r)\beta^r}{\Gamma(\alpha)} \end{aligned}$$

□